

## Theory of EM algorithm

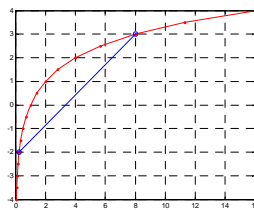
Reference to the book "Pattern Recognition and Machine Learning" by C. M. Bishop.

$$\ln p(\mathbf{X} | \theta) = L(q, \theta) + KL(q || p), \tag{1}$$

Where  $L(q, \theta) = \sum_z q(z) \ln \frac{p(\mathbf{X}, z | \theta)}{q(z)}$  and  $KL(q || p) = -\sum_z q(z) \ln \frac{p(z | \mathbf{X}, \theta)}{q(z)}$

By examining  $KL(q || p)$  which is a Kullback–Leibler divergence of two distributions, we find out  $L(q, \theta)$  is always a lower bound of our objective function  $\ln p(\mathbf{X} | \theta)$ !

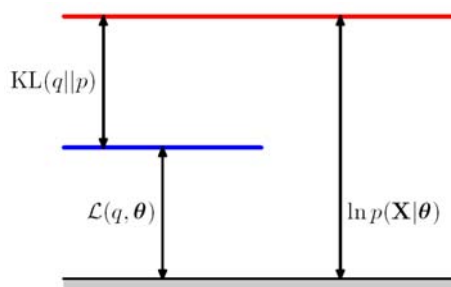
Specifically,  $KL(q || p) \geq 0$  and equality occurs iff  $p = q$ , and we prove it as follows. We first note that  $\ln x$  is a strictly concave function.



i.e.  $\sum_i \alpha_i \ln(x_i) \leq \ln\left(\sum_i \alpha_i x_i\right)$ , for  $\sum_i \alpha_i = 1$ , and equality holds iff  $x_i = 1, \forall i$ . So we

have  $KL(q || p) = -\sum_z q(z) \ln \frac{p(z | \mathbf{X}, \theta)}{q(z)} \geq -\ln\left(\sum_z p(z | \mathbf{X}, \theta)\right) = 0$ , equality hold iff  $p = q$ .

No matter how we chose the distribution  $q(z)$ , (1) always holds. So if we set  $q(z) = p(z | \mathbf{X}, \theta)$ , then we have  $\ln p(\mathbf{X} | \theta) = L(q, \theta)$ .



Let we have a previously estimated parameters  $\theta_{old}$ . From (1) we have

$$\ln p(\mathbf{X} | \theta_{old}) = L(q, \theta_{old}) + KL(q || p) = \sum_z q(z) \ln \frac{p(\mathbf{X}, z | \theta_{old})}{q(z)} + \left( -\sum_z q(z) \ln \frac{p(z | \mathbf{X}, \theta_{old})}{q(z)} \right)$$

If we set  $q^*(z) = p(z | \mathbf{X}, \theta_{old})$ . We have

$$\ln p(\mathbf{X} | \theta_{old}) = L(q^*, \theta_{old}) = \sum_z q^*(z) \ln \frac{p(\mathbf{X}, z | \theta_{old})}{q^*(z)}$$

Setting  $\theta_{new} = \arg \max_{\theta} L(q^*, \theta)$ . We have a new lower bound  $L(q^*, \theta_{new}) \geq L(q^*, \theta_{old})$ .

And we have the following relation:

$$\ln p(\mathbf{X} | \theta_{old}) = L(q^*, \theta_{old}) \leq L(q^*, \theta_{new}) \leq L(q^*, \theta_{new}) + KL(q^* || p) = \ln p(\mathbf{X} | \theta_{new}).$$

$\theta_{new}$  increases the objective function now! Again, setting  $q' = p(z | \mathbf{X}, \theta_{new})$  and maximizing  $L(q', \theta)$ , we could have a sequence of non-decreasing objective function values.

We conclude two main steps here:

(1) E-step: Setting  $q^*(z) = p(z | \mathbf{X}, \theta_{old})$  which induces to the lower bound

$$L(q^*, \theta) = \sum_z q^*(z) \ln \frac{p(\mathbf{X}, z | \theta)}{q^*(z)} \text{ having equality with objective function while}$$

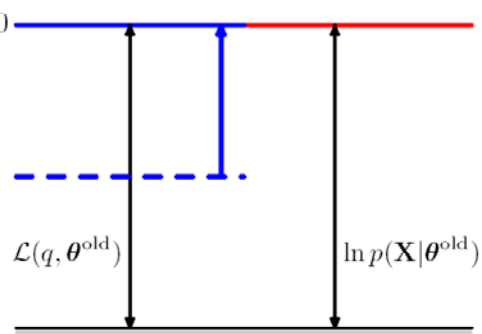
$$\theta = \theta_{old} \text{ (i.e. } \ln p(\mathbf{X} | \theta_{old}) = L(q^*, \theta_{old}) \text{)}.$$

(2) M-step:

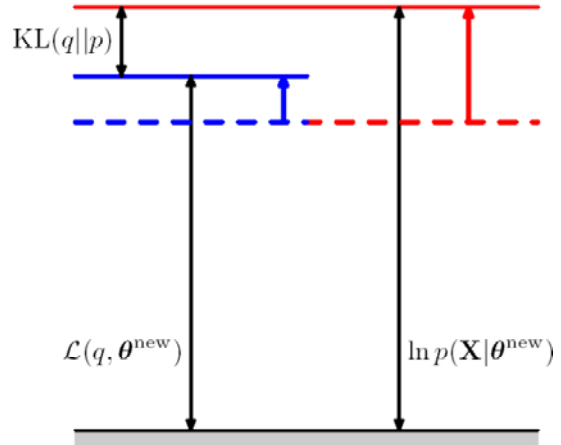
$$\theta_{new} = \arg \max_{\theta} L(q^*, \theta) = \arg \max_{\theta} \sum_z q^*(z) \ln \frac{p(\mathbf{X}, z | \theta)}{q^*(z)} = \arg \max_{\theta} \sum_z q^*(z) \ln p(\mathbf{X}, z | \theta)$$

An alternative interpretation of M-step is finding the parameters that maximize the complete data log-likelihood under the expectation of missing variables.

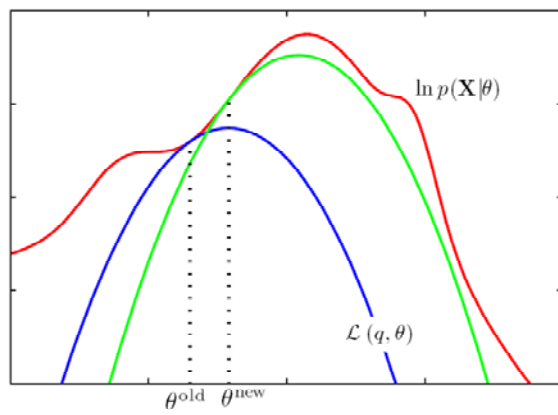
**Figure 9.12** Illustration of the E step of the EM algorithm. The  $q$  distribution is set equal to the posterior distribution for the current parameter values  $\theta^{old}$ , causing the lower bound to move up to the same value as the log likelihood function, with the KL divergence vanishing.



**Figure 9.13** Illustration of the M step of the EM algorithm. The distribution  $q(\mathbf{Z})$  is held fixed and the lower bound  $\mathcal{L}(q, \theta)$  is maximized with respect to the parameter vector  $\theta$  to give a revised value  $\theta^{\text{new}}$ . Because the KL divergence is nonnegative, this causes the log likelihood  $\ln p(\mathbf{X}|\theta)$  to increase by at least as much as the lower bound does.



**Figure 9.14** The EM algorithm involves alternately computing a lower bound on the log likelihood for the current parameter values and then maximizing this bound to obtain the new parameter values. See the text for a full discussion.



Maximum a posteriori could also be achieved by EM. To see this, we note that  $\ln p(\theta | \mathbf{X}) = \ln p(\mathbf{X} | \theta) + \ln p(\theta) - \ln p(\mathbf{X})$ , so we have

$$\begin{aligned} \ln p(\theta | \mathbf{X}) &= \ln p(\mathbf{X} | \theta) + \ln p(\theta) - \ln p(\mathbf{X}) \\ &= L(q, \theta) + \ln p(\theta) + KL(q || p) - \ln p(\mathbf{X}) \end{aligned}$$

The only difference of MAP EM and ML EM is that the M-step is involved maximizing  $L(q, \theta) + \ln p(\theta)$

Also in practice, we usually have multiple observations  $\mathbf{X} = \{x_1 \dots x_N\}$ . The EM

algorithm leads to maximize  $\ln p(\mathbf{X} | \theta) = \ln \prod_i p(x_i | \theta) = \sum_i \ln p(x_i | \theta)$ . So we

adopt EM for each observation  $\ln p(x_i | \theta)$ , and the whole EM cycle is the same as

before except we have a summation for all observations from the outside.